

Estimations

(1) μ and $\gamma(\cdot)$:

prop. If (X_t) is stationary with mean μ

and autocov. $\gamma(\cdot)$. Then:

$$\gamma(k) \xrightarrow{n \rightarrow \infty} 0 \Rightarrow \text{Var}(\bar{X}_n) \xrightarrow{n \rightarrow \infty} 0.$$

Additionally, $\sum_2 |\gamma(k)| < \infty \Rightarrow \text{Var}(\bar{X}_n) \sim \frac{\sum_2 \gamma(k)}{n}$

pf. $\text{Var}(\bar{X}_n) \stackrel{i)}{=} \frac{1}{n} \sum_{|h| < n} (1 - \frac{|h|}{n}) \gamma(h)$

$$\stackrel{ii)}{\leq} \frac{1}{n} \sum_{|h| < n} \gamma(h).$$

Apply stolz on ii) to obtain the first conclusion, the latter is from i).

cor. $(X_t)_2$ is stationary seq. st.

$$X_t = \mu + \sum_2 \gamma_j Z_{t-j}, \quad Z_t \sim \text{IID}(0, \Sigma)$$

If $\sum_2 |\gamma_j| < \infty$, Then:

$$\bar{X}_n \sim AN(\mu, \frac{\sum_2 \gamma(h)}{n})$$

Thm. Under the conditions of cor. above.

If i) $\mathbb{E}(Z_t^p) < \infty$ or ii) $\sum_2 \gamma_j^2 |j| < \infty$.

Then: $\tilde{e}^1(h) \sim AN(e(h), W/n)$.

where $\hat{c}^T(h) = (\hat{c}(1) \dots \hat{c}(h))$. $\hat{c}(h) = \frac{\hat{\gamma}(h)}{\hat{\gamma}(0)}$.

$$W_{ij} = \sum_{k=1}^{\infty} (c(k+i) + c(k-i) - 2c(i)c(k)) (c(k+j) + c(k-j) - 2c(j)c(k)).$$

(2) ARMA models:

Next, we will model $p, q, \vec{\phi}, \vec{\theta}$ and σ^2 .

① AR(p) - process:

1) Consider zero-mean causal AR(p) model:

$$\phi(B)X_t = Z_t, \quad Z_t \sim WN(0, \sigma^2).$$

(By causality: $X_t = Z_t + \sum_{j=1}^{\infty} \psi_j Z_{t-j}$)

product with X_{t-j} , $j=0, \dots, p$. we have:

$$\begin{cases} \sigma^2 = \gamma(0) - \vec{\phi}^T \vec{\gamma}_p, & j=0. \\ \vec{I}_p \vec{\phi} = \vec{\gamma}_p, & j>0. \end{cases} \quad \text{by orthogonality.}$$

We have Yule-Walker estimator:

$$\begin{cases} \hat{\sigma}^2 = \hat{\gamma}(0) - \hat{\vec{\phi}}^T \hat{\vec{\gamma}}_p \\ \hat{\vec{I}}_p \hat{\vec{\phi}} = \hat{\vec{\gamma}}_p. \end{cases} \quad \begin{array}{l} \hat{\vec{I}}_p, \hat{\vec{\gamma}}_p \text{ are estimated} \\ \text{by data } (X_i)_i^{\infty}. \end{array}$$

Then i) $\hat{\vec{\phi}}$ and $\hat{\vec{\gamma}}$ also satisfy:

$$\hat{\gamma}(h) - \hat{\phi}_1 \hat{\gamma}(h-1) - \dots - \hat{\phi}_p \hat{\gamma}(h-p) = \begin{cases} 0 & h=1, \dots, p. \\ \hat{\sigma}^2, & h=0 \end{cases}$$

ii) $n^{\frac{1}{2}}(\hat{\vec{\phi}} - \vec{\phi}) \xrightarrow{d} N(0, \sigma^2 \vec{I}_p^{-1})$. $\hat{\sigma}^2 \xrightarrow{P} \sigma^2$.

ii) p is commonly unknown.

consider for m , use $\hat{\phi}_m = (\hat{\phi}_{m1} \dots \hat{\phi}_{mm})$
 $= \hat{R}_m^{-1} \hat{L}_m$. follows from the relation above.

$$\Rightarrow X_t = \hat{\phi}_{m1} X_{t-1} \dots - \hat{\phi}_{mm} X_{t-m} = Z_t, \text{ where}$$

$$Z_t \sim WN(0, \hat{\gamma}(0) - \hat{\phi}_m^T \hat{\gamma}_p) \triangleq WN(0, V_m)$$

Remark: We can also use the Durbin
 - Levinson algorithm to generate
 $(\hat{\phi}_m)_{ism}$ if $\hat{\gamma}(0) > 0$.

(It also satisfies the relation
 equations)

$$\underline{\text{Thm.}} \quad n^{\frac{1}{2}}(\hat{\phi}_m - \phi_m) \xrightarrow{d} N(0, \sigma^2 I_m).$$

② MA(q)-process:

consider for m , $\hat{\theta}_m = (\hat{\theta}_{m1} \dots \hat{\theta}_{mm})$.

$$X_t = Z_t + \hat{\theta}_{m1} Z_{t-1} + \dots + \hat{\theta}_{mm} Z_{t-m}, \quad Z_t \sim WN(0, \hat{V}_m)$$

We can apply the innovation algorithm to
 obtain the estimators:

$$\hat{\theta}_{m,m-k} = \hat{V}_k^{-1} (\hat{\gamma}(m-k) - \sum_{j=0}^{k-1} \hat{\theta}_{m,m-j} \hat{\theta}_{t,t-j} \hat{V}_j), \quad k=1 \dots m-1.$$

$$\hat{V}_0 = \hat{\gamma}(0) - \sum_{j=0}^{m-1} \hat{\theta}_{m,m-j} \hat{V}_j.$$

Thm. For causal invertible ARMA(p,q) - process

$$\phi(B)X_t = \theta(B)Z_t, \quad Z_t \sim \text{IID}(0, \sigma^2), \quad E(Z_t^4) < \infty.$$

$$\psi(z) = \theta(z) / \phi(z). \quad \text{Then: } \forall (m(n))_{n=1, \dots, \infty} \text{ s.t.}$$

$$m(n) = o(n^{\frac{1}{2}}), \quad (n \rightarrow \infty), \quad \text{we have:}$$

$$n^{\frac{1}{2}} (\hat{\theta}_m - \psi_1, \dots, \hat{\theta}_{mk} - \psi_k)^T \xrightarrow{L} N(0, A)$$

$$\hat{\sigma}_m^2 \xrightarrow{P} \sigma^2, \quad \text{where } a_{ij} = \sum_{k=1}^{j \wedge i} \psi_{i-k} \psi_{j-k}.$$

Remark: For MA(q) - process, $\hat{\theta}_m$ isn't consistent with θ only for special $(m(n))$.

③ ARMA(p,q):

For causal ARMA(p,q) - process:

$$\phi(B)X_t = \theta(B)Z_t, \quad Z_t \sim \text{WN}(0, \sigma^2).$$

$$\Rightarrow X_t = \psi(B)Z_t, \quad \psi(z) = \theta(z) / \phi(z).$$

$$\text{where } \psi_0 = 1, \quad \psi_j = \theta_j + \sum_{i=1}^{j \wedge p} \phi_i \psi_{j-i}, \quad \begin{matrix} \theta_j = 0 & j > q \\ \phi_j = 0 & j > p \end{matrix}$$

By Thm. above, replace ψ_i by $\hat{\theta}_{m,i}$ (approx.).

$$\Rightarrow \hat{\theta}_{m,j} = \theta_j + \sum_{i=1}^{j \wedge p} \phi_i \hat{\theta}_{m,j-i}, \quad j = 1, \dots, p+q.$$

From the equation $j = q+1, \dots, p+q$:

$$\text{We have: } \begin{pmatrix} \hat{\theta}_{m,q+1} \\ \vdots \\ \hat{\theta}_{m,p+q} \end{pmatrix} = \begin{pmatrix} \hat{\theta}_{m,q+1-q} \\ \vdots \\ \hat{\theta}_{m,p+q-p} \end{pmatrix}_{\text{top}} \begin{pmatrix} \phi_1 \\ \vdots \\ \phi_p \end{pmatrix}$$

So, obtain estimate $\hat{\phi}$, replace into equation $j=1 \sim q$.

to get estimate $\hat{\theta}$.